## Proof of Perron-Frobenius Theorem

Let  $x=(x_1,x_2,\ldots,x_n)^T\in\mathbb{R}^n$  with  $y^T$  denoting exclusively the transpose of vector y. Let  $\|x\|=\max_i\{|x_i|\}$  be the norm. Then the induced operator norm for matrix  $A=[a_{ij}]$  is  $\|A\|=\max_i\{\sum_j|a_{ij}|\}$ .

Consider a Markov's chain on n states with transition probabilities  $p_{ij} = \Pr(X_{k+1} = i | X_k = j)$ , independent of k, and  $P = [p_{ij}]$  the transition matrix. Then  $\sum_{i=1}^n p_{ij} = 1$  for all j. Let  $p_{ij}^{(t)} = \Pr(X_{k+t} = i | X_k = j)$  and  $P^{(t)} = [p_{ij}^{(t)}]$  be the t-step transition probability matrix. Then we have  $p_{ij}^{(t)} = \sum_{\ell} p_{i\ell}^{(t-1)} p_{\ell j}$  for all i, j. In matrix,  $P^{(t)} = P^{(t-1)}P = \cdots = P^t$  which is the t-step transition matrix. If  $q = (q_1, \ldots, q_n)^T$  is a probability distribution for the Markovian states at a given iterate with  $q_i \geq 0$ ,  $\sum q_i = 1$ , then Pq is again a probability distribution for the states at the next iterate. A probability distribution w is said to be a steady state distribution if it is invariant under the transition, i.e. Pw = w. Such a distribution must be an eigenvector of P and  $\lambda = 1$  must be the corresponding eigenvalue. The existence as well as the uniqueness of the steady state distribution is guaranteed for a class of Markovian chains by the following theorem due to Perron and Frobenius.

**Theorem 1.** Let  $P = [p_{ij}]$  be a probability transition matrix, i.e.  $p_{ij} \geq 0$  and  $\sum_{i=1}^{n} p_{ij} = 1$  for every j = 1, 2, ..., n. Assume P is irreducible and transitive in the sense that there is a  $t \geq 1$  so that  $p_{ij}^{(t)} > 0$  for all i, j. Then I is a simple eigenvalue of P and all other eigenvalues  $\lambda$  satisfy  $|\lambda| < 1$ . Moreover, the unique eigenvector can be chosen to be a probability vector w and it satisfies  $\lim_{t\to\infty} P^t = [w, w, \ldots, w]$ . Furthermore, for any probability vector q we have  $P^t q \to w$  as  $t \to \infty$ .

*Proof.* We first prove a claim that  $\lim_{t\to\infty} p_{ij}^{(t)}$  exist for all i,j and the limit is independent of j,  $\lim_{t\to\infty} p_{ij}^{(t)} = w_i$ .

Because  $P = [p_{ij}]$  (is irreducible and transitive) has non-zero entries, we have

$$\delta = \min_{ij} p_{ij} > 0.$$

Consider the equation of the ijth entry of  $P^{t+1} = [p_{ij}^{(t+1)}] = P^t P$ ,

$$p_{ij}^{(t+1)} = \sum_{k} p_{ik}^{(t)} p_{kj}.$$

Let

$$0 < m_i^{(t)} := \min_j p_{ij}^{(t)} \leq \max_j p_{ij}^{(t)} := M_i^{(t)} < 1.$$

Then, we have

$$m_i^{(t+1)} = \min_j \sum_k p_{ik}^{(t)} p_{kj} \ge m_i^{(t)} \sum_k p_{kj} = m_i^{(t)}.$$

i.e., the sequence  $\{m_i^{(1)}, m_i^{(2)}, \dots\}$  is non-decreasing. Similarly, the upper bound sequence  $\{M_i^{(1)}, M_i^{(2)}, \dots\}$  is non-increasing. As a result, both limits  $\lim_{t \to \infty} m_i^{(t)} = m_i \leq M_i = \lim_{t \to \infty} M_i^{(t)}$  exist. We now prove they are equal  $m_i = M_i$ .

To this end, we consider the difference  $M_i^{(t+1)} - m_i^{(t+1)}$ :

$$M_{i}^{(t+1)} - m_{i}^{(t+1)} = \max_{j} \sum_{k} p_{ik}^{(t)} p_{kj} - \min_{\ell} \sum_{k} p_{ik}^{(t)} p_{k\ell}$$

$$= \max_{j,\ell} \sum_{k} p_{ik}^{(t)} (p_{kj} - p_{k\ell})$$

$$= \max_{j,\ell} \left[ \sum_{k}^{+} p_{ik}^{(t)} (p_{kj} - p_{k\ell}) + \sum_{k}^{-} p_{ik}^{(t)} (p_{kj} - p_{k\ell}) \right]$$

$$\leq \max_{j,\ell} \left[ M_{i}^{(t)} \sum_{k}^{+} (p_{kj} - p_{k\ell}) + m_{i}^{(t)} \sum_{k}^{-} (p_{kj} - p_{k\ell}) \right]$$
(1)

where  $\sum_{k}^{+} p_{ik}^{(t)}(p_{kj} - p_{k\ell})$  means the summation of all non-negative terms  $p_{kj} - p_{k\ell} \geq 0$  and similarly  $\sum_{k}^{-} p_{ik}^{(t)}(p_{kj} - p_{k\ell})$  means the summation of all negative terms  $p_{kj} - p_{k\ell} < 0$ .

It is critical to notice the following unexpected equality:

$$\sum_{k}^{-} (p_{kj} - p_{k\ell}) = \sum_{k}^{-} p_{kj} - \sum_{k}^{-} p_{k\ell}$$

$$= 1 - \sum_{k}^{+} p_{kj} - (1 - \sum_{k}^{+} p_{k\ell})$$

$$= \sum_{k}^{+} (p_{k\ell} - p_{kj})$$

$$= -\sum_{k}^{+} (p_{kj} - p_{k\ell}).$$

Hence, the inequality (1) becomes

$$M_i^{(t+1)} - m_i^{(t+1)} \le (M_i^{(t)} - m_i^{(t)}) \max_{j,\ell} \sum_k^+ (p_{kj} - p_{k\ell}).$$

If  $\max_{j,\ell} \sum_k^+ (p_{kj} - p_{k\ell}) = 0$ , which is independent of all t, it is done that  $M_i^{(t)} = m_i^{(t)}$ . Otherwise, for the pair  $j,\ell$  that gives the maximum let r be the number of terms in k for which  $p_{kj} - p_{k\ell} > 0$ , and s be the number of terms for which  $p_{kj} - p_{k\ell} < 0$ . Then  $r \ge 1$ , and  $\tilde{n} := r + s \ge 1$  as well as  $\tilde{n} \le n$ . More importantly

$$\sum_{k}^{+} (p_{kj} - p_{k\ell}) = \sum_{k}^{+} p_{kj} - \sum_{k}^{+} p_{k\ell}$$

$$= 1 - \sum_{k}^{-} p_{kj} - \sum_{k}^{+} p_{k\ell}$$

$$\leq 1 - s\delta - r\delta = 1 - \tilde{n}\delta$$

$$\leq 1 - \delta < 1.$$

The estimate for the difference  $M_i^{(t+1)} - m_i^{(t+1)}$  at last reduces to

$$M_i^{(t+1)} - m_i^{(t+1)} \le (1 - \delta)(M_i^{(t)} - m_i^{(t)}) \le (1 - \delta)^t (M_i^{(1)} - m_i^{(1)}) \to 0,$$

as  $t \to \infty$ , showing  $M_i = m_i := w_i$ . As a consequence to the inequality  $m_i^{(t)} \le p_{ij}^{(t)} \le M_i^{(t)}$ , we have  $\lim_{t \to \infty} p_{ij}^{(t)} = w_i$  for all j. In matrix notation,  $\lim_{t \to \infty} P^t = [w, w, \dots, w] := W$ , a matrix of equal column vectors.

Next, assume there is a  $k \ge 1$  so that only  $p_{ij}^{(k)} > 0$  for all i, j. Then the result above implies  $P^{kt} \to W$  as  $t \to \infty$ . We need to show that  $P^t \to W$  as  $t \to \infty$  as well. This is left as an exercise.

Next, we show that  $\lambda=1$  is an eigenvalue with eigenvector w. In fact from the definition of w above  $\lim_{t\to\infty}P^t=[w,w,\ldots,w]$  and thus  $[w,w,\ldots,w]=\lim_{t\to\infty}P^t=P\lim_{t\to\infty}P^{t-1}=P[w,w,\ldots,w]=[Pw,Pw,\ldots,Pw]$  showing Pw=w.

To show that the eigenvalue  $\lambda=1$  is simple, two cases are considered. First, let  $x\neq 0$  be an eigenvector Px=x of the eigenvalue 1. Apply P to the identity repeatedly to have  $P^tx=x$ . In limit,  $\lim_{t\to\infty}P^tx=Wx=[w,w,\dots,w]x=wx_1+wx_2+\dots+wx_n=(\sum x_j)w=x$ . Denote it by  $\bar x=\sum_j x_j$ . Then  $x_i=\bar xw_i$  for all i. Because  $x\neq 0$ , we must have  $\bar x=\sum_j x_j\neq 0$ , and that  $x=\bar xw$  for some constant  $\bar x\neq 0$ , showing that the eigenvector w is unique up to a constant multiple. Second, let  $x\neq 0$  be a generalized eigenvector of  $\lambda=1$ . Then there is a constant  $c\neq 0$  so that Px=x+cw which implies  $P^tx=x+ctw$ . Since  $\lim_{t\to\infty}P^tx=[w,w,\dots,w]x$  exists on the left, we must have c=0 on the right, a contradiction. Together we can conclude that the dimension of the generalized eigenspace for  $\lambda=1$  is 1, i.e., the eigenvalue 1 is simple. In addition, for any probability vector q, the result above shows  $\lim_{t\to\infty}P^tq=Wq=\sum_j q_jw=w$  as  $Wx=\sum_j x_jw$  always holds.

Next, let  $\lambda$  be an eigenvalue of P. Then it is also an eigenvalue for the transpose  $P^T$ . Let x be an eigenvector of  $\lambda$  of  $P^T$ . Then  $P^Tx = \lambda x$  and  $\|\lambda x\| = |\lambda| \|x\| = \|P^Tx\| \le \|P^T\| \|x\|$ . Since  $\|P^T\| = 1$  because  $\sum_{i=1}^n p_{ij} = 1$  we have  $|\lambda| \le 1$ .

Next, let x be an eigenvector of an eigenvalue  $\lambda$ . Then we have  $\lim_{t\to\infty} P^t x = Wx = (\sum x_j)w$  on one hand and  $\lim_{t\to\infty} P^t x = \lim_{t\to\infty} \lambda^t x$  on the other hand. So either  $|\lambda| < 1$  in which case  $\lim_{t\to\infty} \lambda^t x = 0$  and then  $\sum x_j = 0$ , or  $|\lambda| = 1$  in which case  $\lambda = e^{i\theta}$  for some  $\theta$  and the limit  $\lim_{t\to\infty} \lambda^t = \lim_{t\to\infty} e^{i\theta t}$  exists since  $\lim_{t\to\infty} e^{i\theta t} x = \lim_{t\to\infty} \lambda^t x = \lim_{t\to\infty} P^t x = (\sum x_j)w$ . The latter case holds if and only if  $\sum x_j \neq 0$  and  $\theta = 0$ , i.e.,  $\lambda = 1$ . This shows that all eigenvalues that is not  $\lambda = 1$  are inside the unit circle and the corresponding eigenspace is  $\{x: \sum_j x_j = 0\}$ , which n-1 dimensional.  $\square$ 

References: Bellman(1997); Berman & Plemmons(1994); Frobenius(1908, 1912); Lancaster & Tismenetsky(1985); Marcus & Minc(1984); Perron(1907); Petersen(1983); Seneta(1973).

Ethier and Kurtz, Markov Processes – Characterization and Convergence.