

Math Finance Seminar: Numerical Simulation of SDEs

T. Shores

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Outline

- 1 Brownian Motion
- 2 Stochastic Integrals
- 3 Stochastic Differential Equations
- 4 Euler-Maruyama Method
- 5 Convergence of EM Method
- 6 Milstein's Higher Order Method
- 7 Linear Stability
- 8 Stochastic Chain Rule
- 9 Parting Shots

References

- ① Desmond Higham, *An Algorithmic Introduction to Numerical Simulation of Stochastic Differential Equations*, Siam Rev. **43(3)**, 2001, p. 525–546.
- ② Peter Kloeden and Eckhard Platen, *Numerical Solution of Stochastic Differential Equations*, Springer, New York, 1999.
- ③ Robert Hogg and Allen Craig, *Introduction to Mathematical Statistics*, 5th Ed., Prentice-Hall, Englewood, N. J., 1995.

Standard (continuous) Brownian motion, or Wiener process, over $[0, T]$: a random variable $W(t)$ depending continuously on $t \in [0, T]$ such that

- 1 $W(0) = 0$ with probability 1.
- 2 For $0 \leq s < t \leq T$ the random variable

$$W(t) - W(s) \sim N(0, t - s).$$

- 3 For $0 \leq s < t < u < v \leq T$ the random variables $W(t) - W(s)$ and $W(v) - W(u)$ are independent.

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Discretized Brownian Motion

Discretized Brownian motion over $[0, T]$ in N steps: a sequence of random variable $W_j = W(t_j)$, where $\delta t = T/N$ and $t_j = j \delta t$, such that

- 1 $W(0) = 0$ with probability 1.
- 2 For $j = 1, 2, \dots, N$, $W_j = W_j + dW_j$.
- 3 For $j = 1, 2, \dots, N$, $dW_j \sim N(0, \delta t)$.

Notice that items (1)–(3) of continuous Brownian motion follow from these conditions. In fact, thanks to independence and identical distributions,

$$W_{j+k} - W_j = \sum_{i=1}^k dW_i \sim N(0, k \delta t).$$

(Recall, for independent X, Y , $E[aX + bY] = aE[X] + bE[Y]$ and $\text{var}(aX + bY) = a^2 \text{var}(X) + b^2 \text{var}(Y)$.)

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Here is the file used by Higham. Let's run it and play with the parameters. In particular, rem out the resetting of the random number generator:

```
% BPATH2 Brownian path simulation:  vectorized
randn('state',100) % set the state of randn
T = 1; N = 500; dt = T/N;
dW = sqrt(dt)*randn(1,N); % increments
W = cumsum(dW); % cumulative sum
plot([0:dt:T],[0,W],'r-') % plot W against t
xlabel('t','FontSize',16)
ylabel('W(t)','FontSize',16,'Rotation',0)
```

Function of Brownian Motion Simulation

We can also simulate random walks that are functions of Brownian motion. Here is the example of

$$X(t) = u(W(t), t) = e^{(t + \frac{1}{2}W(t))}$$

```
%BPATH3 Function along a Brownian path
randn('state',100) % set the state of randn
T = 1; N = 500; dt = T/N; t = [dt:dt:1];
M = 1000; % M paths simultaneously
dW = sqrt(dt)*randn(M,N); % increments
W = cumsum(dW,2); % cumulative sum
U = exp(repmat(t,[M 1]) + 0.5*W);
Umean = mean(U);
plot([0,t],[1,Umean],'b-'), hold on % plot mean over M
paths
plot([0,t],[ones(5,1),U(1:5,:)],'r--'), hold off % plot 5
individual paths
xlabel('t','FontSize',16)
ylabel('U(t)','FontSize',16,'Rotation',0,'HorizontalAlignment','right')
legend('mean of 1000 paths','5 individual paths',2)
averr = norm((Umean - exp(9*t/8)),'inf') % sample error
```

Ito

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$$X(t) - X(0) = \int_0^t h(\tau) dW(\tau)$$

provided that $X(t)$ is a random process such that

$$X(t) - X(0) = \lim_{m \rightarrow \infty} \sum_{j=0}^{N-1} h(t_j) (W(t_{j+1}) - W(t_j))$$

where $0 = t_0 < t_1 < \dots < t_N = t$ and $\max_j (t_{j+1} - t_j) \rightarrow 0$ as $N \rightarrow \infty$.

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A Special Case: Ito

Take $h(t) = W(t)$. Some shorthand: $W_j = W(t_j)$,
 $W_{j+1/2} = W(t_j + \frac{\delta t}{2}) = W(t_j + t_{j+1})$ and $dW_j = W_{j+1} - W_j$.
 Thus $W_N = W(T)$ and $W_0 = W(0)$. For the Ito integral:

- Note the identity

$$b(a - b) = \frac{1}{2} (a^2 - b^2 - (a - b)^2)$$

- Hence

$$\begin{aligned}
 \sum_{j=0}^{N-1} W_j (W_{j+1} - W_j) &= \frac{1}{2} \sum_{j=0}^{N-1} (W_{j+1}^2 - W_j^2 - (dW_j)^2) \\
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- Now recall from statistics that i.i.d. r.v.'s $X_1, \dots, X_N \sim N(\mu, \sigma^2)$, then

$$Y = \sum_{j=1}^N \left(\frac{X_j - \mu}{\sigma} \right)^2 \sim \chi^2(N),$$

which has mean N and variance $2N$.

- Hence, since $\delta W_j \sim N(0, \delta t)$, we have that

$$\sum_{j=0}^{N-1} (dW_j)^2 = \sum_{j=0}^{N-1} (W_{j+1} - W_j)^2 \sim \delta t \chi^2(N).$$

- Thus, this sum has mean $N \delta t = T$ and variance $\delta t^2 2N = 2T \delta t$.
- So it is reasonable that the sum approaches T as $\delta t \rightarrow 0$.
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 W_{j+1/2} &= \frac{W_j + W_{j+1}}{2} + \frac{1}{2} (W_{j+1/2} - W_{j+1}) + \frac{1}{2} (W_{j+1/2} - W_j) \\
 &= \frac{W_j + W_{j+1}}{2} + \frac{1}{2} (-U_j) + \frac{1}{2} (V_j),
 \end{aligned}$$

where $U_j, V_j \sim N(0, \frac{\delta t}{2})$ are independent r.v.'s.

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- Now expand (at the board) and really get the telescoping effect, so the sum becomes

$$\begin{aligned}\sum_{j=0}^{N-1} W_{j+1/2} (W_{j+1} - W_j) &= \sum_{j=0}^{N-1} \left(\frac{W_j + W_{j+1}}{2} + \Delta Z_j \right) (W_{j+1} - W_j) \\ &= \frac{1}{2} (W(T)^2 - W(0)^2) + \sum_{j=0}^{N-1} \Delta Z_j (W_{j+1} - W_j)\end{aligned}$$

- Each term in the latter sum is a $\frac{1}{2} (V_j^2 - U_j^2)$, so has mean zero and variance $\frac{\delta t^2}{4}$, since $U_j^2, V_j^2 \sim \frac{\delta t}{2} \chi^2(1)$ are independent.
- Hence, the sum of these independent variables is a random variable of mean zero and variance $N \delta t \frac{\delta t}{4} = \frac{T}{4} \delta t$.
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Simulations

The file stint.m:

```
% Ito and Stratonovich integrals of W dW
randn('state',100) % set the state of randn
T = 1; N = 500; dt = T/N;
dW = sqrt(dt)*randn(1,N); % increments
W = cumsum(dW); % cumulative sum
ito = sum([0,W(1:end-1)].*dW)
strat = sum((0.5*([0,W(1:end-1)]+W) +
0.5*sqrt(dt)*randn(1,N)).*dW)
itoerr = abs(ito - 0.5*(W(end)^2-T))
straterr = abs(strat - 0.5*W(end)^2)
```

Deterministic Definitions

Deterministic Differential Equation:

To compute a function $x(t)$, $0 \leq t \leq T$, such that on the interval $[0, T]$, given $x(0)$ (this is an IVP, really):

- Derivative form: $\frac{dx}{dt} = f(x, t)$.
- Differential form: $dx = f(x, t) dt$.
- Integral form: $x(t) = x(0) + \int_0^t f(x(s), s) ds$.
- Each has a point of view about the ODE, but these are all equivalent definitions involving deterministic variability $f(x, t)$.

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Stochastic Differential Equation:

To compute a stochastic process $X(t)$, $0 \leq t \leq T$, such that on the interval $[0, T]$, given $X(0)$ (this is an IVP, really):

- We not only want to account for deterministic variability, $f(X(t), t)$, but also stochastic variability:

- Differential form:

$$dX(t) = f(X(t), t) dt + g(X(t), t) dW(t).$$

- Integral form:

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- Caution: either form forces us to make a choice about which is the appropriate stochastic integral to use.

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$$dX(t) = f(X(t), t) dt + g(X(t), t) dW(t).$$

- Integral form:

$$X(t) = X(0) + \int_0^t f(X(s), s) ds + \int_0^t g(X(s), s) dW(s).$$

- Caution: either form forces us to make a choice about which is the appropriate stochastic integral to use.

Stochastic Definitions

Stochastic Differential Equation:

To compute a stochastic process $X(t)$, $0 \leq t \leq T$, such that on the interval $[0, T]$, given $X(0)$ (this is an IVP, really):

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Example (Risky Asset Pricing, a.k.a., Geometric Brownian Motion):

An asset price $X(t)$ can be viewed as a random process. The relative change in price, $dX(t)/X(t)$ can be viewed as having two (additive) components:

- A deterministic factor: λdt . If there were no risk, we could think of λ as the growth rate over time. In the simplest case, λ is constant.
- A random factor: $\mu dW(t)$, where $dW = \sqrt{dt}Z$, $Z \sim N(0, 1)$ and $W(t)$ is Brownian motion. In the simplest case, μ is constant.
- So the stochastic differential equation that results is the linear differential equation

$$\frac{dX(t)}{X(t)} = \lambda dt + \mu dW(t)$$

or $dX(t) = \lambda X(t) dt + \mu X(t) dW(t)$ (multiplicative noise).

- Exact solution: $X(t) = X(0) e^{(\lambda - \frac{1}{2}\mu^2)t + \mu W(t)}$.

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Deterministic Case

Numerical Solutions:

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Convergence and stability of the Euler methods:

- Classical analysis shows that under reasonable conditions, the methods are convergent of order one in Δt , i.e.,
$$\| [x_j - x(t_j)] \| = \mathcal{O}(\Delta t), \delta t \rightarrow 0.$$
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 $X_{j+1} = X_j + f(X_j, \tau_j) \Delta t + g(X_j, \tau_j) (W(\tau_{j+1}) - W(\tau_j))$
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Computational Example

Implementation Convention: A discrete Brownian path is generated using dt . Then the Euler-Maruyama time step is a multiple of dt , say $R * dt = \Delta t$.

```
%EM Euler-Maruyama method on linear SDE
%
% SDE is  $dX = \lambda X dt + \mu X dW$ ,  $X(0) = X_{\text{zero}}$ ,
% where  $\lambda = 2$ ,  $\mu = 1$  and  $X_{\text{zero}} = 1$ .
%
% Discretized Brownian path over  $[0,1]$  has  $dt = 2^{-8}$ .
% Euler-Maruyama uses timestep  $R*dt$ .
randn('state',100)
lambda = 2; mu = 1; Xzero = 1; % problem parameters
T = 1; N = 2^8; dt = T/N;
dW = sqrt(dt)*randn(1,N); % Brownian increments
W = cumsum(dW); % discretized Brownian path
```

Computational Example Continued

```
Xtrue = Xzero*exp((lambda-0.5*mu^2)*([dt:dt:T])+mu*W);
plot([0:dt:T],[Xzero,Xtrue],'m-'), hold on
R = 4; Dt = R*dt; L = N/R; % L EM steps of size Dt = R*dt
Xem = zeros(1,L); % preallocate for efficiency
Xtemp = Xzero;
for j = 1:L
    Winc = sum(dW(R*(j-1)+1:R*j));
    Xtemp = Xtemp + Dt*lambda*Xtemp + mu*Xtemp*Winc;
    Xem(j) = Xtemp;
end
plot([0:Dt:T],[Xzero,Xem],'r--*'), hold off
xlabel('t','FontSize',12)
ylabel('X','FontSize',16,'Rotation',0,'HorizontalAlignment','right')
emerr = abs(Xem(end)-Xtrue(end))
```

Numerical Method for $dX = f(X, t) dt + g(X, t) dW$ on $[0, T]$:

- Converges **strongly** if mean of the error converges to zero, i.e.,

$$\lim_{n \rightarrow \infty} E [|X_n - X(\tau)|] = 0,$$

- and with **order of convergence** γ if there exists $C > 0$ such that for any fixed $\tau = n \Delta t \in [0, T]$,

$$E [|X_n - X(\tau)|] \leq C \Delta t^\gamma$$

for all Δt sufficiently small. Put another way, the expected value of the error is $\mathcal{O}(\Delta t)$, $\Delta t \rightarrow 0$.

- Uniform order convergence does follow for EM, but this isn't obvious, nor is it the form of the definition of strong convergence in Kloeden-Platen, as is apparently the case here.

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An Experiment

Idea Behind the Experiment:

- If you think that there is a valid order condition

$$E_{\Delta t} \leq C \Delta t^\gamma,$$

- assume that the inequality is sharp and replace it by

$$E_{\Delta t} \approx C \Delta t^\gamma.$$

- Take logs of both sides and get

$$Y_{\Delta t} = \log E_{\Delta t} \approx \log C + \gamma \log \Delta t.$$

- Do a log-log plot of $E_{\Delta t}$ against Δt .
- A graph that resembles a straight line of slope γ and intercept $\log C$ supports your suspicion.

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An Experiment Continued

Compute geometric Brownian motion by taking the mean of 1000 different Brownian paths on $[0, 1]$ at $T = \tau = 1$. Use $\delta t = 2^{-9}$ and $\Delta t = 2^{p-1} \delta t$, $1 \leq p \leq 5$. Then do a log-log plot, linear regression to estimate γ (q in the program), and the norm of the residual:

```
%EMSTRONG Test strong convergence of Euler-Maruyama
% Solves  $dX = \lambda X dt + \mu X dW$ ,  $X(0) = X_{\text{zero}}$ ,
% where  $\lambda = 2$ ,  $\mu = 1$  and  $X_{\text{zero}} = 1$ .
% Discretized Brownian path over  $[0, 1]$  has  $dt = 2^{-9}$ .
% E-M uses 5 different timesteps: 16dt, 8dt, 4dt, 2dt, dt.
% Examine strong convergence at  $T=1$ :  $E | X_L - X(T) |$ .
randn('state',100)
lambda = 2; mu = 1; Xzero = 1; % problem parameters
T = 1; N = 2^9; dt = T/N; %
M = 1000; % number of paths sampled
Xerr = zeros(M,5); % preallocate array
for s = 1:M, % sample over discrete Brownian paths
    dW = sqrt(dt)*randn(1,N); % Brownian increments
    W = cumsum(dW); % discrete Brownian path
    Xtrue = Xzero*exp((lambda-0.5*mu^2)+mu*W(end));
```

An Experiment Continued

```
for p = 1:5
R = 2^(p-1); Dt = R*dt; L = N/R; % L Euler steps of size Dt =
R*dt
Xtemp = Xzero;
for j = 1:L
Winc = sum(dW(R*(j-1)+1:R*j));
Xtemp = Xtemp + Dt*lambda*Xtemp + mu*Xtemp*Winc;
end
Xerr(s,p) = abs(Xtemp - Xtrue); % store the error at t = 1
end
end
Dtvals = dt*(2.^([0:4]));
subplot(221) % top LH picture
loglog(Dtvals,mean(Xerr),'b*-'), hold on
loglog(Dtvals,(Dtvals.^(.5)),'r--'), hold off % reference slope
of 1/2
axis([1e-3 1e-1 1e-4 1])
xlabel('\Delta t'), ylabel('Sample average of | X(T) - X_L |')
title('emstrong.m','FontSize',10)
%%% Least squares fit of error = C * Dt^q %%%
A = [ones(5,1), log(Dtvals)']; rhs = log(mean(Xerr)');
sol = A\rhs; q = sol(2)
resid = norm(A*sol - rhs)
```

Numerical Method for

$dX(t) = f(X(t), t) dt + g(X(t), t) dW(t)$ on $[0, T]$:

- Converges **weakly** if mean of functions of the error taken from some set of test functions (like polynomials, which would give moments) converges to zero, i.e.,

$$\lim_{n \rightarrow \infty} |E[p(X_n)] - E[p(X(\tau))]| = 0,$$

- and with **order of convergence** γ if there exists $C > 0$ such that for any fixed $\tau = n \Delta t \in [0, T]$,

$$|E[p(X_n)] - E[p(X(\tau))]| \leq C \Delta t^\gamma$$

for all Δt sufficiently small.

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$$|E[p(X_n)] - E[p(X(\tau))]| \leq C \Delta t^\gamma$$

for all Δt sufficiently small.

An Experiment

Note: We have assumed that errors other than sampling error like floating point error and sampling bias are negligible compared to sampling error. This is reasonable in relatively small experiments.

```
%EMWEAK Test weak convergence of Euler-Maruyama
% Solves  $dX = \lambda X dt + \mu X dW$ ,  $X(0) = X_{\text{zero}}$ ,
% where  $\lambda = 2$ ,  $\mu = 1$  and  $X_{\text{zero}} = 1$ .
% E-M uses 5 different timesteps:  $2^{-(p-10)}$ ,  $p = 1, 2, 3, 4, 5$ .
% Examine weak convergence at  $T=1$ :  $|E(X_L) - E(X(T))|$ .
% Different paths are used for each E-M timestep.
% Code is vectorized over paths.
% Uncommenting the line indicated below gives the weak E-M
method.
randn('state',100);
lambda = 2; mu = 0.1; Xzero = 1; T = 1; % problem parameters
M = 50000; % number of paths sampled
Xem = zeros(5,1); % preallocate arrays
for p = 1:5 % take various Euler timesteps
Dt =  $2^{-(p-10)}$ ; L = T/Dt; % L Euler steps of size Dt
Xtemp = Xzero*ones(M,1);
```

An Experiment Continued

```
for j = 1:L
Winc = sqrt(Dt)*randn(M,1);
% Winc = sqrt(Dt)*sign(randn(M,1)); %% use for weak E-M %%
Xtemp = Xtemp + Dt*lambda*Xtemp + mu*Xtemp.*Winc;
end
Xem(p) = mean(Xtemp);
end
Xerr = abs(Xem - exp(lambda));
Dtvals = 2.^([1:5]-10);
subplot(222) % top RH picture
loglog(Dtvals,Xerr,'b*-'), hold on
loglog(Dtvals,Dtvals,'r--'), hold off % reference slope of 1
axis([1e-3 1e-1 1e-4 1])
xlabel('\Delta t'), ylabel('| E(X(T)) - Sample average of X_L |')
title('emweak.m','FontSize',10)
%%% Least squares fit of error = C * dt^q %%%
A = [ones(p,1), log(Dtvals)']; rhs = log(Xerr);
sol = A\rhs; q = sol(2)
resid = norm(A*sol - rhs)
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The Method

A careful study of Ito-Taylor expansions leads to a higher order method (Milstein's method):

$$X_{j+1} = X_j + f(X_j, \tau_j) \Delta t + g(X_j, \tau_j) (W(\tau_{j+1}) - W(\tau_j)) \\ + \frac{1}{2} g(X_j) g_x(X_j, \tau_j) \left((W(\tau_{j+1}) - W(\tau_j))^2 - \Delta t \right)$$

An Experiment

Now run the experiment `milstrong.m` to solve the population dynamics stochastic differential equation (the stochastic Verhulst equation)

$$dX(t) = rX(t)(K - X(t))dt + \beta X(t)dW(t)$$

which is simply a stochastic logistic equation.

One interesting aspect of the program: the exact (strong) solution is well known, but involves another stochastic integral. Hence, the most accurate solution (smallest Δt) is used as a “reference” solution.

The Deterministic Case

Long Term Stability of the Euler Methods:

- Does *not* mean stability on finite intervals, which would require that perturbations in initial conditions cause perturbations in the computed solution that remain bounded as $\delta t \rightarrow 0$.
- Define the **linear stability domain** of a method to be the subset $D = \{z = \lambda \Delta t \mid \lim_{j \rightarrow \infty} x_j = 0\}$ where the sequence $\{x_j\}$ is produced by applying the method to the model problem $dx/dt = \lambda x$, $x(0) = 1$.
- The method is **A-stable** if D contains the open left half-plane. Reason: negative $\Re(\lambda)$ and positive Δt are main parameters of interest for this asymptotic (or absolute) stability.

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Long Term Stability in Stochastic Setting:

- The model problem is

$$dX(t) = \lambda X(t) dt + \mu X(t) dW(t).$$

Solution:

$$X(t) = X(0) e^{(\lambda - \frac{1}{2}\mu^2)t + \mu W(t)}$$

- The mathematical stability of a solution comes in two flavors, assuming that $X(0) \neq 0$ with probability 1.
- Mean-square stability:

$$\lim_{t \rightarrow \infty} E[X(t)^2] = 0 \iff \Re(\lambda) + \frac{1}{2}|\mu|^2 < 0.$$

- Stochastic asymptotic stability:

$$\lim_{t \rightarrow \infty} |X(t)|^2 = 0, \text{ with probability 1 } \iff \Re\left(\lambda - \frac{1}{2}\mu^2\right) < 0.$$

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Long Term Stability of Numerical Method:

One can show:

- Mean-square stability of a numerical method:

$$\lim_{j \rightarrow \infty} E[X_j^2] = 0 \iff |1 + \Delta t \lambda|^2 + \frac{1}{2} \Delta t |\mu|^2 < 0.$$

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$$\lim_{j \rightarrow \infty} |X_j^2| = 0, \text{ with probability 1,}$$

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$$\begin{aligned} \lim_{j \rightarrow \infty} |X_j^2| &= 0, \text{ with probability 1,} \\ &\iff E \left[\log \left| 1 + \Delta t \lambda + \sqrt{\Delta t} \mu N(0, 1) \right| \right] < 0. \end{aligned}$$

Run the script `stab.m`. Settings are $\Delta t = 1, 1/2, 1/4$, $\lambda = 1/2$, and $\mu = \sqrt{6}$. For asymptotic stability, run over a single path, while for mean-square stability, an average of paths. Note, ideally in mean-square case we should have straight line graphs, since we calculate logy graphs.

Deterministic Case

Let's start with the deterministic chain rule: given a function $F(x, t)$, the first order differential is given by

$$df = \frac{\partial F(x, t)}{\partial x} dx + \frac{\partial F(x, t)}{\partial t} dt,$$

which gives first order (linear) approximations by the Taylor formula. Of course, if $x = x(t)$, we simply plug that into the formula for the one variable differential. We might reason accordingly that if $X = X(t)$, is a stochastic process, then we should be able to plug X into x and get the correct differential. Wrong! Well, at least if you use Ito integrals. (With Stratonovich integrals you would be right.)

Stochastic Chain Rule

For a function $F(X, t)$ of a stochastic process $X(t)$:

- Start over with a Taylor expansion

$$\begin{aligned} dF = & \frac{\partial F(x, t)}{\partial x} dx + \frac{\partial F(x, t)}{\partial t} dt + \frac{1}{2} \frac{\partial^2 F(x, t)}{\partial x^2} dx^2 \\ & + \frac{\partial^2 F(x, t)}{\partial x \partial t} dx dt + \frac{1}{2} \frac{\partial^2 F(x, t)}{\partial t^2} dt^2. \end{aligned}$$

- Now make the substitutions $x = X(t)$ and $dx = dX(t) = f(X(t), t) dt + g(X(t), t) dW(t)$.
- For a first order (linear) approximation, we have no problem in discarding the higher order dt^2 term.
- Nor does the mixed term present a problem:

$$dX dt = (f(X(t), t) dt + g(X(t), t) dW(t)) dt.$$

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Ito formula for $dX(t) = f(X(t), t)dt + g(X(t), t)dW(t)$:

- The problem is with the second order term in dx^2 because $dW^2 \sim \delta t \chi^2(1)$, which has mean δt and variance $2\delta t^2$. So it is reasonable that the term approaches δt as $\delta t \rightarrow 0$.
- The net result is that

$$dF = \frac{\partial F(X, t)}{\partial X} dX + \frac{\partial F(X, t)}{\partial t} dt + \frac{1}{2} \frac{\partial^2 F(X, t)}{\partial X^2} dX^2.$$

- Substitute $dX = f dt + g dW$, discard $dW dt$ and dt^2 terms and get

$$dF = \left(F_X f + F_t + \frac{1}{2} F_{XX} g^2 \right) dt + F_X g dW$$

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Example

The linear model for volatile stock price $X(t)$ with drift λ and volatility μ

$$dX(t) = \lambda X(t) dt + \mu X(t) dW(t).$$

Suppose a portfolio consists of an option (buy or sell) for a share of the stock with price $p(X, t)$, and a short position of Δ shares of it. It's value: $F = p(X, t) - \Delta X$. By the Ito formula,

$$dF = \left((p_X - \Delta) \lambda X + p_t + \frac{1}{2} p_{XX} \mu^2 X^2 \right) dt + (p_X - \Delta) \mu X dW.$$

Some Elementary Inequalities

Theorem

(Markov Inequality) Let $u(X)$ be a non-negative function of the r.v. X with finite expected value. For all positive a ,

$$P(u(X) \geq a) \leq \frac{E[u(X)]}{a}.$$

Theorem

(Chebychev Inequality) If the r.v. X has finite variance σ^2 and expected value μ , then for all positive k

$$P[|X - \mu| \geq k\sigma] \leq \frac{1}{k^2}.$$

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Remarkably simple!

- For Markov, let event $E = \{x \mid u(x) \geq a\}$ and $f(x)$ be p.d.f. of X .

$$E[u(X)] = \int_{-\infty}^{\infty} u(x) f(x) dx \geq \int_E a f(x) dx = aP(E).$$

Now divide by a and we're done!

- For Chebychev, take $u(X) = (X - \mu)^2$, $a = \sigma^2 k^2$ and obtain from Markov

$$P(|X - \mu| \geq k\sigma) \equiv P((X - \mu)^2 \geq \sigma^2 k^2) \leq \frac{\sigma^2}{\sigma^2 k^2} = \frac{1}{k^2}.$$

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Done!

Applications

Recall that a sequence X_n of r.v.'s **converges in probability** to r.v. X if for all $\epsilon > 0$,

$$\lim_{n \rightarrow \infty} P(|X_n - X| \geq \epsilon) = 0.$$

With this definition and the previous theorems, we can explain “it is reasonable that...”

“It is reasonable that...”

- (In derivation for $\int_0^T W(t) dW(t)$,
“ $\sum_{j=0}^{N-1} (dW_j)^2 = \sum_{j=0}^{N-1} (W_{j+1} - W_j)^2 \sim \delta t \chi^2(N)$. Thus,
this sum has mean $N \delta t = T$ and variance $\delta t^2 2N = 2T \delta t$. So
it is reasonable that the sum approaches T as $\delta t \rightarrow 0$.”
- Let $X_N = \sum_{j=0}^{N-1} (dW_j)^2$ and let $k = 1/\sqrt{\delta t}$, so that

$$k\sigma = \frac{1}{\sqrt{\delta t}} 2T \delta t = 2T \frac{\sqrt{T}}{\sqrt{N}} = \frac{2T^{3/2}}{\sqrt{N}}, \text{ and } k^2 = \frac{1}{\delta t} = \frac{N}{T}.$$

- Hence

$$P \left[|X_N - T| \geq \frac{2T^{3/2}}{\sqrt{N}} \right] \leq \frac{T}{N} \rightarrow 0 \text{ as } N \rightarrow \infty.$$

- Hence X_N converges to T in probability.

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One More Application

- Suppose that an iterative method is strongly convergent at $\tau = T$ of order γ , so that

$$E [|X_n - X(\tau)|] \leq C \Delta t^\gamma.$$

- It follows that

$$\frac{E [|X_n - X(\tau)|]}{\Delta t^{\gamma/2}} \leq C \Delta t^{\gamma/2}.$$

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